

Graph Theory-Based Network Analysis of Regional Uncertainties of the US Economy

Rangan Gupta

Department of Economics, University of Pretoria, Pretoria, 0002, South Africa. Email:

rangan.gupta@up.ac.za.

Chi-Keung (Marco) Lau*

Huddersfield Business School, University of Huddersfield, Huddersfield, HD1 3DH, United

Kingdom. Email: c.lau@hud.ac.uk.

Xin Sheng

Lord Ashcroft International Business School, Anglia Ruskin University, Chelmsford, CM1

1SQ, United Kingdom. Email: xin.sheng@anglia.ac.uk.

Abstract

We study the transmission mechanism of time-varying macroeconomic uncertainty across the US states. We analyse the contemporaneous and temporal causal relationships of uncertainty at the state level by utilising the Bayesian graphical VAR (BGVAR) model. Our results show that the current uncertainty of each state strongly depends on the previous level of uncertainty in its own state. We find evidence of strong contemporaneous and lagged dependence among US states. The analysis of this paper has important policy implications.

JEL Codes: C32

Keywords: Uncertainty Spillover, Macroeconomic Uncertainty, US States, Bayesian Graphical Structural VAR

* Corresponding author.

1. Introduction

The effect of uncertainty on macroeconomy and financial markets has attracted great attention of financial academics and practitioners in the wake of the 2008 global financial crisis (e.g., Antonakakis, Balcilar, Gupta, and Kyei, 2017; Bloom, 2009, Chuliá, Gupta, Uribe, and Wohar, 2017, Gupta, Hammoudeh, Modise, and Nguyen, 2014; among others). A number of recent studies (e.g., Angelini et al., 2018; Carriero et al., 2016; Jurado et al., 2015; Mumtaz, 2018; among others) find evidence that macroeconomic uncertainty is a driver of economic fluctuations and it plays an important role in affecting real economic activity and asset pricing. While there is a growing interest in understanding uncertainty and its impact on macroeconomic and financial variables, the transmission mechanism of uncertainty across borders remains an underexplored area. The research on uncertainty interconnectedness has important implications for policy responses to the global financial crisis. If uncertainty spills over across borders, then even there is no change of uncertainty in the domestic market, the negative impact of uncertainty can still present itself through trade and financial linkages among economies (Klößner and Sekkel, 2014; Balli, Uddin, Mudassar, and Yoon, 2017). Furthermore, uncertainty spillover effects can magnify the adverse influence when domestic uncertainty does rise (Bernal, Gnabo, and Guilmin, 2016).¹ The few existing studies focus mainly on international uncertainty linkages at the country level and provide some evidence that uncertainty transmits across national borders (e.g., Antonakakis, Gabauer, and Gupta, 2018;

¹Klößner and Sekkel (2014) and Balli et al. (2017) provide some economic explanations about the transmission channels through which uncertainty spillovers take place from the perspectives of the trade and financial linkages and the negative impact of economic policy uncertainty on macroeconomic fundamentals. For example, the effect of increasing policy uncertainty in one economy may affect economic fundamentals (such as capital flows, bond risk premia, etc.) of other economies, with the potential effect of rising economic policy uncertainty in these economies. Baker, Bloom and Davis (2016) find evidence of adverse effects of economic policy uncertainty on investment and employment on the firm and macroeconomic levels. Bernal et al. (2016) and Bai, Zhang, Liu, Wang (2019) find empirical evidence that uncertainty leads to high degrees of risk spillovers and negatively affects real economy activities and financial markets in other economies.

Antonakakis, Gabauer, Gupta, Plakandaras, 2018; Christou, Gozgor, Gupta, and Lau, 2019; Gabauer and Gupta, 2018; Gupta, Pierdzioch, and Risse 2016; Gupta, Lau, and Wohar, 2016; Huang, Tong, Qiu, and Shen, 2018; Klößner and Sekkel, 2014; Yin and Han, 2014;). Our contribution to this emerging stream of literature is to examine the behaviour of macroeconomic uncertainty dependence among US states using the state-level data. If there is a linkage of uncertainty across state economies, the changes of macroeconomic uncertainty in one state can have an impact on economic and financial variables of other states even there are no changes in their domestic levels of uncertainty of these states. This paper contributes to economic policy-making decisions of state governors. It is of great value for policy makers to understand the transmission channel of uncertainty spillovers given the important role of uncertainty in affecting macroeconomy (Balcilar et al., 2016). In the US, states have sufficiently autonomy in policies they adopted and state governors bear responsibility for the performance of the state economy (Brown, 2010).

The state-level uncertainty data used in this paper is based on the h-step-ahead forecast of a factor-augmented vector autoregression (FAVAR) system estimated by Mumtaz (2018) using the uncertainty measure proposed by Jurado et al. (2015). The new measure provides direct econometric estimates of time-varying uncertainty under a data-rich environment, which is free from both the restriction of specific theoretical models and the dependence of any individual macroeconomic variables.² The Bayesian Graphical Structural VAR (BGVAR) model of Ahelegbey et al. (2016) is employed to examine the contemporaneous and dynamic causal structures of macroeconomic uncertainty across the 50 US states. This newly developed model

² Existing studies (e.g., Zhang, Lei, Ji, and Kutan, 2019; Gabauer and Gupta, 2018; among others) mainly use news-based measures of economic policy uncertainty (EPU) to quantify measures of uncertainty. This paper uses the uncertainty measure of Jurado et al. (2015) to reflect the comprehensive uncertainty of macroeconomic fundamentals of state economies.

is superior to the standard Structural VAR (SVAR) model that is often criticised for imposing implausible structural restrictions based on a specific economic theory. The BGVAR model allows for the investigation of causal relationships between variables without the restriction of economic models.³ Moreover, it provides a framework to represent and estimate an unambiguous direction of causations by means of the directed edges.⁴ This model adapts an approach where important variables can be identified with a causal interpretation. Ahelegbey et al. (2016) confirm that the BGVAR methodology is more parsimonious and offers a better representation of the causal relations among variables than the Granger causality approach.⁵

³ This is an appealing feature of the model employed in this paper due to the lack of underlying economic theory for the emerging macroeconomic uncertainty interconnectedness literature. A number of studies (e.g., Ji, Bouri, Roubaud, and Kristoufek, 2019; Ji, Bouri, Lau, and Roubaud, 2019; Luo and Ji, 2018; Zhang, Lei, Ji, and Kutan, 2019; among others) employ the total spillover index of Diebold and Yilmaz (2014) to measure interconnectedness. However, as pointed out by Antonakakis, Gabauer, Gupta, and Plakandaras (2018), the methodology arbitrarily sets the rolling window-size and there is a loss of observations/information in the process

⁴ Ahelegbey et al. (2016) show that one of the appealing features of the BGVAR model is the possibility of giving a graphical representation of the logical implications of models. A directed acyclic graph (DAG) (i.e., the edges of DAG are directed and connected without circles so that these edges can only flow forward and the graph is not cyclic) is used in this approach to represent an unambiguous direction of causal relationship between variables. For example, the relationship $A \rightarrow B$ means that the variable A causes the variable B. The node A (ancestor) from which a directed edge originates is the explanatory variable, and the node B (descendant) to which the directed edge ends is the response variable. If $A \rightarrow B \rightarrow C$, A and C would be probabilistically dependent in the absence of B. The edge probabilities (i.e., the posterior probabilities of the presence of edges) are produced by the model under the MAR structure (i.e., the multivariate autoregression structure, which captures the dynamic causal relationship between variables and detects edges that are persistent over time) and the MIN structure (i.e., multivariate instantaneous structure, which presents the contemporaneous dependence among variables).

⁵ Ahelegbey et al. (2016) suggest that the BGVAR model offers a more accurate representation of the linkages among variables than the Granger causality approach. The traditional pairwise Granger causality test (P-GC) only deals with bivariate time series and does not consider the conditioning on relevant covariates, and the modified conditional Granger causality test (C-PC) has a problem of over-parametrisation which leads to inefficiency in accurately gauging the causal relationships.

The remainder of this paper is organised as follow. Section 2 describes the data and methodology. Section 3 presents the results and analysis. Section 4 provides a summary of findings and concluding remarks.

2. Data and Methodology

The US state-level uncertainty data used in this paper is based on the h-step-ahead ($h = 1, 2, 3, 4$) forecast of a factor-augmented vector autoregression (FAVAR) system estimated by Mumtaz (2018). The dataset consists of quarterly macroeconomic uncertainty measures for the 50 US states at four different forecast horizons (i.e. in 3, 6, 9, and 12 months) over the period from 1977:Q2 to 2015:Q3. The state-level macroeconomic uncertainty measures are calculated using the real per-capita growth rates of personal income, benefit income, dividend income, social insurance contributions, other income, the seasonally adjusted employment growth rate, unemployment rate, and house price growth rate. These time series are obtained from the Federal Reserve Bank of St Louis database. The uncertainty measures for each state are constructed following the procedures proposed by Jurado et al. (2015).⁶

In order to examine the contemporaneous and lagged causal relationships of macroeconomic uncertainty across US states, this paper utilises the Bayesian graphical VAR (BGVAR) model of Ahelegbey et al. (2016). The Dynamic Bayesian Network is applied to the following standard structural VAR (SVAR) model presented in equation (1).⁷

⁶ See online technical appendix of Mumtaz (2018) for details of the uncertainty construction procedures.

⁷ See Dagum et al. (1992) for details of the Dynamic Bayesian Network technique.

$$Y_t = B_0 Y_t + \sum_{i=1}^p B_i Y_{t-i} + \varepsilon_t \quad (1)$$

where Y_t is a vector of response variables. p is the lag order. $\varepsilon_t \sim N(0, I_p)$. B_0 is a $(n_y \times n_y)$ matrix of structural parameters with zeros on the diagonal.

The SVAR model of Eq (1) can be written into the reduced form of VAR as follows:

$$Y_t = \sum_{i=1}^p A_i Y_{t-i} + u_t \quad (2)$$

where $A_i = A_0^{-1} B_i^*$, $1 \leq i \leq p$. $A_0 = (I - B_0)$. $u_t = A_0^{-1} \varepsilon_t$.

It is noteworthy that A_0 are not identified, which require some identification restrictions to perform structural analysis. The SVAR model is thus widely criticised for imposing implausible assumptions or, at least, assumptions that are only as credible as the underlying economic models (Ahelegbey et al., 2016). This critique motivates the use of the Bayesian graphical VAR (BGVAR) model of Ahelegbey et al. (2016) in this paper, in light of the lack of underlying economic theory for the emerging macroeconomic uncertainty spillovers literature. The BGVAR model has two distinctive advantages over the SVAR model. First, it is not necessary to impose restrictions from an economic theory in the BGVAR model to identify the causal order of structural models. Second, the BGSVAR model offers insight into the contemporaneous and dynamic (temporal) dependence of response variables with a causal interpretation, and it provides a simple framework to represent and estimate an unambiguous direction of causation among the variables by means of the directed edges. There is a one-to-one relationship between the regression matrices of the SVAR model and a directed acyclic graph (DAG), given as:

$$X_{t-s}^j \rightarrow X_t^i \Leftrightarrow B_{s,ij}^* \neq 0, \quad 0 \leq s \leq p \quad (3)$$

where X_t^i represents the realisation value of the i -th variable at time t . The relationship $X_{t-s}^j \rightarrow X_t^i$ means that X_{t-s}^j causes X_t^i . It can be referred to as contemporaneous causal relationships for $s = 0$, and as lagged dependence for $1 \leq s \leq p$.

$$\text{Define } B_s^* = (G_s \circ \Phi_s), \quad 0 \leq s \leq p \quad (4)$$

where G_s is a binary connectivity matrix that indicates dependence, and Φ_s is a coefficient matrix. The operator \circ is the Hadamard product. G_0 represents the connectivity matrix of contemporaneous dependence. G_s ($1 \leq s \leq p$) denotes the connectivity matrix of the temporal dependence.

There is a one-to-one correspondence between regression matrices and the directed acyclic graphs such that.

$$B_{s,ij}^* = \begin{cases} \phi_{s,ij} & \text{if } B_{s,ij}^* = 1 \\ 0 & \text{if } B_{s,ij}^* = 0 \end{cases} \quad (5)$$

Based on the SVAR in Eq (1), the DAG can be represented as follows

$$Y_t = (G_0 \circ \Phi_0)Y_t + \sum_{i=1}^p (G_i \circ \Phi_i) Y_{t-i} + \varepsilon_t \quad (6)$$

where $(G_i \circ \Phi_i)$ are the graphical model structural coefficient matrices whose non-zero elements describe the value associates with the instantaneous and lagged dependences.

The estimation of the model requires the choice of the optimal lag order, a set of parameters, $\{B_0^*, B_1^*, \dots, B_p^*, \Sigma_\varepsilon\}$, and the inference of causal structure $G = (G_0, G_1, \dots, G_p)$.⁸ The model produces the posterior probabilities of edges for both temporal and contemporaneous relationships, namely multivariate autoregressive (MAR) and multivariate instantaneous (MIN) structures.

3. Results and Analysis

Tables 1-2 summarise the results of edge probabilities for 50 US states of both MAR and MIN structures. Tables 1 reports the dynamic (lagged) dependence of the MAR structure (for $h = 1$), and Tables 2 presents the contemporaneous relationship of the MIN structure (for $h = 1$).⁹ Tables 3-4 show the number of cases that US states are the origination of directed edges for the MAR and MIN structures based on macroeconomic uncertainty measures at four different forecast horizons. The results reveal the following causality patterns based on the posterior probability of 0.50 or above.

In the case of MAR structure, we find that the current uncertainty of each state strongly depends on the previous level of uncertainty in its own state. In Table 1, we find that Arizona, California, Iowa, Idaho, Kansas, Louisiana, Missouri, North Carolina, New Hampshire, New York, and Ohio are important transmitters of uncertainty to other states.¹⁰ For example, the lagged uncertainty level in California is more likely to explain current uncertainty in nine states, i.e.,

⁸ See Ahelegbey et al. (2016) for details of the statistical inference and estimation procedures. The optimal lag order is set to 1 based on the Bayesian Information Criterion (BIC), and 50,000 draws are used.

⁹ The results of edge probabilities for 50 US states of both MAR and MIN structures (for $h = 2, 3, 4$) are available upon request.

¹⁰ These states are highlighted in bold in the tables.

California_{t-1} → (California_t, Georgia_t, Maryland_t, Minnesota_t, Ohio_t, Oklahoma_t, Pennsylvania_t, South Carolina_t, and Vermont_t), with a probability higher than 0.50. In contrast, the current level of uncertainty in California only strongly depends on the previous level of uncertainty in its own state and in Arizona i.e., (California_{t-1} Arizona_{t-1}) → California_t. The results at horizons h1 to h4 are consistent to some extent. For example, Table 3 shows strong evidence of directed edges originating from California, Iowa, Kansas, Louisiana, and New Hampshire to at least nine states at all four horizons. Moreover, we find that preceding uncertainty in Alabama, District of Columbia, Delaware, Hawaii, Michigan, North Dakota, and Oregon also provides important information in explaining the structural dynamics of uncertainty in a relatively large number of states. There are at least nine cases that these states show strong evidence of directed edges from explanatory to response states. The results show that the dynamic spillover effect (temporal dependence) of uncertainty among US states is strong at all four forecast horizons.

Tables 2 report results of the contemporaneous interconnectedness among US states for the MIN structure. The results show that the origination of a large number of direct edges concentrates on a relatively small number of states. In Table 2, there is strong evidence of the contemporaneous causality from Massachusetts to fifteen states (i.e., Arizona, California, Florida, Kentucky, Maine, North Carolina, North Dakota, Nebraska, New Mexico, Nevada, New York, Pennsylvania, Utah, Wisconsin, West Virginia) at horizon h1, i.e., current uncertainty of these states strongly depends on the current level of uncertainty in Tennessee at the h1 horizon. Table 4 presents the number of cases that US states are the origination of directed edges for the MIN structure. MIN reveals strong contemporaneous causations originated from Alaska, Kansas, Kentucky, Idaho, Maryland, Massachusetts, Montana, Nevada, North Carolina, Texas, and Vermont at horizon h1. There are at least nine cases that these states

are the origination of directed edges with a probability higher than 0.50. The results show evidence of strong contemporaneous dependence across all four forecast horizons.

The analysis has important policy implications in that it highlights the transmission channel of regional uncertainty of the US Economy. A major implication is that policymakers should not neglect uncertainty spillovers from other states when making policy decisions. The changes of macroeconomic uncertainty of one state can transmit across state borders and affect economic and financial variables (such as, personal income, dividend income, employment growth rate, unemployment rate, house price growth rate, and so on) of other states.

4. Conclusions

The Bayesian graphical VAR (BGVAR) model is employed in this paper to study the contemporaneous and dynamic lagged dependence of macroeconomic uncertainty across 50 US states. This study examines the behaviour of macroeconomic uncertainty interconnectedness using the state-level data and finds evidence of strong contemporaneous and temporal causal relationships among US states. The results show that the current uncertainty of each state strongly depends on the previous level of uncertainty in its own state. Moreover, the results of the contemporaneous interconnectedness show that the origination of direct edges concentrates on a few states, i.e., current uncertainty in a large number of states strongly depends on the current level of uncertainty in a relatively small number of states. The findings of this paper has important implications for policy-making decisions.

As part of future research, it might be interesting to conduct rolling estimation analysis to examine the behaviour of dependence over time, especially during the recent crisis period. It might also be worthwhile to explore the potential determinants of cross-state uncertainty spillovers and to provide theoretical and economic explanations of the observed transmission channels.

References

- Ahelegbey D. F., Billio, M. and Casarin, R. (2016), Bayesian Graphical Models for Structural Vector Autoregressive Processes. *Journal of Applied Econometrics*, 31(2), 357–386.
- Angelini, G., Bacchiocchi, E., Caggiano, G., & Fanelli, L. (2018). Uncertainty across volatility regimes. CESifo Working Paper Series No. 6799
- Antonakakis, N., Balcilar, M., Gupta, R., & Kyei, C. (2017). Components of economic policy uncertainty and predictability of US stock returns and volatility: evidence from a nonparametric causality-in-quantile approach. *Frontiers in Finance and Economics*, 14(2), 20-49.
- Antonakakis, N., Gabauer, D., and Gupta, R. (2018). Greek Economic Policy Uncertainty: Does it Matter for the European Union? University of Pretoria, Department of Economics, Working Paper No. 201840.
- Antonakakis, N., Gabauer, D., Gupta, R., and Plakandaras, V. (2018). Dynamic connectedness of uncertainty across developed economies: A time-varying approach. *Economics Letters*, 166, 63–75.
- Bai, L., Zhang, X., Liu, Y., & Wang, Q. (2019). Economic risk contagion among major economies: New evidence from EPU spillover analysis in time and frequency domains. *Physica A: Statistical Mechanics and its Applications*, 535, 122431.

Baker, S. R., Bloom, N., & Davis, S. J. (2016). Measuring economic policy uncertainty. *The quarterly journal of economics*, 131(4), 1593-1636

Balli, F., Uddin, G. S., Mudassar, H., & Yoon, S. M. (2017). Cross-country determinants of economic policy uncertainty spillovers. *Economics Letters*, 156, 179-183.

Bernal, O., Gnabo, J. Y., & Guilmin, G. (2016). Economic policy uncertainty and risk spillovers in the Eurozone. *Journal of International Money and Finance*, 65, 24-45.

Bloom, N. (2009). The impact of uncertainty shocks. *econometrica*, 77(3), 623-685.

Brown, A. (2010). Are Governors Responsible for the State Economy? Partisanship, Blame, and Divided Federalism. *The Journal of Politics*, 72(3), 605-615.

Mumtaz, H. (2018). Does uncertainty affect real activity? Evidence from state-level data. *Economics Letters*, 167,

Christou, C., Gozgor, G., Gupta, R., & Lau, C. K. (2019). Are Uncertainties across the World Convergent? University of Pretoria, Department of Economics, Working Paper No. 201907.

Chuliá, H., Gupta, R., Uribe, J. M., & Wohar, M. E. (2017). Impact of US uncertainties on emerging and mature markets: Evidence from a quantile-vector autoregressive approach. *Journal of International Financial Markets, Institutions and Money*, 48, 178-191.

Dagum, P., Galper, A., & Horvitz, E. (1992, July). Dynamic network models for forecasting. In *Proceedings of the eighth international conference on uncertainty in artificial intelligence* (pp. 41-48). Morgan Kaufmann Publishers Inc..

Gabauer, D. and Gupta, R. (2018). On the transmission mechanism of country-specific and international economic uncertainty spillovers: Evidence from a TVP-VAR connectedness decomposition approach. *Economics Letters*, 171, 63–71.

Gupta, R., Lau, C. K. M., & Wohar, M. E. (2016). The impact of US uncertainty on the Euro area in good and bad times: evidence from a quantile structural vector autoregressive model. *Empirica*, 1-16.

- Gupta, R., Hammoudeh, S., Modise, M. P., & Nguyen, D. K. (2014). Can economic uncertainty, financial stress and consumer sentiments predict US equity premium?. *Journal of International Financial Markets, Institutions and Money*, 33, 367-378.
- Gupta, R., Pierdzioch, C., & Risse, M. (2016). On international uncertainty links: BART-based empirical evidence for Canada. *Economics Letters*, 143, 24-27.
- Huang, Z., Tong, C., Qiu, H., & Shen, Y. (2018). The spillover of macroeconomic uncertainty between the US and China. *Economics Letters*, 171, 123-127.
- Jurado, K., Ludvigson, S. C., & Ng, S. (2015). Measuring uncertainty. *American Economic Review*, 105(3), 1177-1216.
- Ji, Q., Bouri, E., Roubaud, D., & Kristoufek, L. (2019). Information interdependence among energy, cryptocurrency and major commodity markets. *Energy Economics*, 81, 1042-1055.
- Ji, Q., Bouri, E., Lau, C. K. M., & Roubaud, D. (2019). Dynamic connectedness and integration in cryptocurrency markets. *International Review of Financial Analysis*, 63, 257-272.
- Jiang, Y., Zhu, Z., Tian, G., & Nie, H. (2019). Determinants of within and cross-country economic policy uncertainty spillovers: Evidence from US and China. *Finance Research Letters*.
- Klößner, S., & Sekkel, R. (2014). International spillovers of policy uncertainty. *Economics Letters*, 124(3), 508-512.
- Luo, J., & Ji, Q. (2018). High-frequency volatility connectedness between the US crude oil market and China's agricultural commodity markets. *Energy Economics*, 76, 424-438.
- Yin, L., & Han, L. (2014). Spillovers of macroeconomic uncertainty among major economies. *Applied Economics Letters*, 21(13), 938-944.
- Zhang, D., Lei, L., Ji, Q., & Kutan, A. M. (2019). Economic policy uncertainty in the US and China and their impact on the global markets. *Economic Modelling*, 79, 47-56.

Table 1 Results of the MAR structure at the h1 horizon

Response States_ t	Explanatory States_ t-1										
Alaska	Alaska	Nebraska	Pennsylvania	West Virginia							
	1.00	0.50	0.52	0.68							
Alabama	Alabama	Arizona	Missouri	North Carolina	South Carolina	South Dakota					
	1.00	0.56	0.79	0.57	0.73	0.50					
Arkansas	Arkansas	Colorado	Massachusetts	New Hampshire							
	0.97	0.81	0.51	0.58							
Arizona	Alabama	Arizona	North Carolina	Ohio	South Dakota						
	0.51	0.98	0.78	0.55	0.58						
California	Arizona	California									
	0.53	1.00									
Colorado	Colorado	District of Columbia	Kansas	Nebraska	New Mexico						
	1.00	0.54	0.57	0.60	0.51						
Connecticut	Connecticut	Illinois	Kansas	Kentucky	Massachusetts	North Dakota	New Mexico	Nevada	Rhode Island	Vermont	Washington
	1.00	0.60	0.55	0.53	0.60	0.58	0.85	0.56	0.54	0.50	0.59
District of Columbia	District of Columbia	Hawaii	Idaho	Nevada							
	1.00	0.50	0.51	0.51							
Delaware	Arkansas	Colorado	Delaware	North Carolina	North Dakota	New York					
	0.52	0.53	1.00	0.59	0.50	0.66					
Florida	Alabama	Florida	Iowa	Louisiana	Maine	Missouri	New Hampshire	Ohio	South Dakota		

	0.57	0.81	0.52	0.54	0.76	0.52	0.57	0.79	0.52
Response States_ t	Explanatory States_ t-1								
Georgia	California	District of Columbia	Georgia	Missouri	New Hampshire	Ohio	Washington		
	0.50	0.59	0.98	0.94	0.50	0.60	0.96		
Hawaii	Colorado	District of Columbia	Hawaii	Idaho	New Hampshire	Ohio			
	0.50	0.56	1.00	0.78	0.56	0.51			
Iowa	Iowa	Idaho	Illinois	Michigan	New Mexico	Utah			
	0.99	0.62	0.64	0.55	0.56	0.56			
Idaho	Idaho	Illinois	Louisiana						
	1.00	0.55	0.79						
Illinois	Connecticut	Idaho	Illinois	Mississippi	Nevada	New York	Virginia		
	0.55	0.53	1.00	0.55	0.63	0.89	0.51		
Indiana	Iowa	Illinois	Indiana	Kansas	New Hampshire	Utah			
	0.82	0.53	1.00	0.50	0.51	0.53			
Kansas	Iowa	Kansas	Montana	New York	Oregon				
	0.69	1.00	0.50	0.51	0.68				
Kentucky	Alabama	Kentucky	Maine	North Carolina	New Jersey	New York	Ohio	Wisconsin	
	0.51	0.96	0.60	0.51	0.51	0.63	0.60	0.60	
Louisiana	Idaho	Louisiana	Maine	North Carolina	Rhode Island				
	0.82	1.00	0.63	0.59	0.68				
Massachusetts	Alaska	Connecticut	Massachusetts	Maine	North Carolina	Ohio	Rhode Island	Virginia	Wisconsin
	0.55	0.59	1.00	0.64	0.53	0.51	0.52	0.59	0.83
Maryland	California	Idaho	Louisiana	Maryland	New Hampshire	New Mexico	Washington	Wisconsin	
	0.53	0.55	0.83	0.99	0.54	0.50	0.60	0.59	
Maine	Idaho	Louisiana	Maine	North Carolina	North Dakota	Rhode Island			

	0.53	0.56	0.50	0.56	0.59	0.99	0.54	0.92	
Response States_ t	Explanatory States_ t-1								
New York	Kansas 0.64	Missouri 0.50	New York 1.00	West Virginia 0.53					
Ohio	California 0.54	District of Columbia 0.52	Idaho 0.51	Kansas 0.53	Louisiana 0.63	New Hampshire 0.57	Ohio 1.00	Utah 0.50	
Oklahoma	California 0.56	Iowa 0.51	Montana 0.59	New York 0.81	Oklahoma 1.00				
Oregon	Hawaii 0.53	Iowa 0.77	Idaho 0.55	Kansas 0.70	Massachusetts 0.52	New Hampshire 0.52	Oregon 1.00		
Pennsylvania	Alaska 0.55	Arizona 0.52	California 0.60	Connecticut 0.51	Kansas 0.72	Ohio 0.53	Pennsylvania 0.94	South Dakota 0.60	West Virginia 0.85
Rhode Island	Kansas 0.67	Maine 0.52	Missouri 0.70	North Carolina 0.91	Oklahoma 0.54	Rhode Island 1.00			
South Carolina	Arizona 0.57	California 0.57	Idaho 0.50	Louisiana 0.50	Hampshire 0.51	South Carolina 0.92	Utah 0.59		
South Dakota	District of Columbia 0.57	Georgia 0.57	Kansas 0.51	Missouri 0.63	North Carolina 0.62	Ohio 0.55	South Dakota 0.90	Tennessee 0.52	Washingto n 0.70
Tennessee	Arizona 0.51	Iowa 0.61	Mississippi 0.53	Hampshire 0.59	New Jersey 0.51				
Texas	Alabama 0.50	Kansas 0.75	Texas 0.99						
Utah	Iowa 0.88	New Mexico 0.67	Utah 1.00						
Virginia	Louisiana	Massachusetts	Maine	Oklahoma	South Carolina	Utah	Virginia		

	0.52	0.80	0.69	0.50	0.56	0.54	1.00
Response States_ t	Explanatory States_ t-1						
Vermont	California	Iowa	Illinois	Mississippi	New Hampshire	Oregon	Vermont
	0.64	0.67	0.50	0.59	0.66	0.57	0.81
Washington	Maryland	Nevada	Washington				
	0.50	0.89	1.00				
Wisconsin	Alaska	Alabama	Kentucky	Oklahoma	South Dakota	Wisconsin	
	0.63	0.63	0.59	0.51	0.50	1.00	
West Virginia	Alaska	Kansas	New York	Ohio	Pennsylvania	South Dakota	West Virginia
	0.61	0.56	0.56	0.51	0.55	0.57	0.99

This table summarises results of the selected edges for the MAR structure based on posterior probabilities greater than 0.5. The states highlighted indicate strong evidence of directed edges originating from explanatory to response states. There are at least nine cases that explanatory states have strong evidence of directed edges to response states.

Table 2 Results of the MIN structure at the h1 horizon

Response States_ t	Explanatory States_ t									
Alaska	Null									
Alabama	Kentucky	Maine	Montana							
	0.90	0.77	0.58							
Arkansas	Maryland	Vermont								
	0.52	0.50								
Arizona	Alaska	Colorado	Kentucky	Mississippi	North Carolina	Nevada	South Carolina	Vermont		
	0.53	0.51	0.55	0.66	0.97	0.80	0.50	0.56		
California	Alaska	Massachusetts	Maryland	Minnesota	North Carolina					
	0.99	0.56	0.86	0.65	0.50					
Colorado	Arkansas	Kansas	Maine	Minnesota	Nevada	Tennessee	Texas			
	0.61	0.63	0.58	0.71	0.53	0.57	0.89			
Connecticut	Idaho	Illinois	Massachusetts	North Carolina	New Mexico					
	0.60	0.76	0.96	0.67	0.75					
District of Columbia	Colorado	Mississippi	Nevada	Ohio						
	0.58	0.72	0.92	0.53						
Delaware	Arkansas	Colorado	Indiana	Kansas	Michigan	New Mexico	New York	Rhode Island	Wisconsin	
	0.75	0.58	0.80	0.56	0.50	0.58	0.53	0.68	0.63	
Florida	Alaska	Alabama	Indiana	Kentucky	Maine	Nevada	Ohio	West Virginia		
	0.87	0.58	0.63	0.55	0.82	0.54	0.57	0.58		
Georgia	Nevada	Texas	Vermont	Washington						
	0.79	0.51	0.91	0.51						
Hawaii	Colorado	Idaho	Maryland	North Carolina	Ohio	Vermont				
	1.00	0.94	0.55	0.55	0.77	0.56				

Response States_ t	Explanatory States_ t							
Iowa	Idaho 0.62	Illinois 0.65	Indiana 0.58	Kansas 0.73	Minnesota 0.57	Mississippi 0.54	Montana 0.64	Vermont 0.67
Idaho	Massachusetts 0.51	Maryland 0.58	Montana 0.73					
Illinois	Maryland 0.52	Montana 0.60						
Indiana	Idaho 0.52	Illinois 0.54	Kansas 0.87	Massachusetts 0.66	North Carolina 0.65	Tennessee 0.60	Texas 0.68	Vermont 0.55
Kansas	Montana 0.62	Texas 0.61						
Kentucky	Alaska 0.66							
Louisiana	Connecticut 0.70	Idaho 0.97	Maryland 0.76	Maine 0.74	Rhode Island 0.58	Washington 0.51		
Massachusetts	Null							
Maryland	Massachusetts 0.55							
Maine	Alaska 0.75	Idaho 0.73	Kentucky 0.74	Massachusetts 0.80	Texas 0.50			
Michigan	Idaho 0.58	Massachusetts 1.00	Mississippi 0.58	Montana 0.73	New Mexico 0.81	Texas 0.77		
Minnesota	Idaho 0.55	Maryland 0.57						
Missouri	Alabama 0.69	Georgia 0.54	Iowa 0.54	Louisiana 0.74	Mississippi 0.85	New Mexico 0.94	Vermont 0.51	Washington 0.70
Mississippi	Illinois 0.70	Massachusetts 0.77	Maryland 0.79	Vermont 0.74				

Response States_ t	Explanatory States_ t								
Montana	Null								
North Carolina	Alaska	Kentucky	Massachusetts	Maine	Montana				
	0.60	0.53	0.86	0.53	0.90				
North Dakota	Alaska	Connecticut	Idaho	Kentucky	Maine	North Carolina			
	0.60	0.60	0.52	0.53	0.63	0.62			
Nebraska	Alaska	Alabama	Kansas	Mississippi	North Carolina	Vermont			
	0.53	0.71	0.60	0.79	0.74	0.62			
New Hampshire	Colorado	Indiana	Kansas	Massachusetts	Maryland	Minnesota	Montana	Nebraska	
	0.68	0.59	0.89	0.61	0.86	0.77	0.71	0.56	
New Jersey	Arkansas	Colorado	Illinois	Indiana	Minnesota	Nevada	Tennessee		
	0.70	0.81	0.80	0.51	0.93	0.63	0.76		
New Mexico	Alaska	Kansas	Massachusetts	Maryland					
	0.75	0.63	0.94	0.69					
Nevada	Alaska	Illinois	Kentucky	Mississippi	Tennessee				
	0.53	0.62	0.57	0.50	0.51				
New York	Alaska	Alabama	Illinois	Kentucky	Montana	Nevada	West Virginia		
	0.55	0.63	0.65	0.54	0.92	0.50	0.85		
Ohio	Idaho	Kansas	Kentucky	Montana	Texas				
	0.82	0.73	0.64	0.54	0.80				
Oklahoma	California	Kansas	Kentucky	Minnesota	Montana	Tennessee	Texas	Vermont	Wisconsin
	0.50	0.50	0.63	0.54	0.95	0.55	0.68	0.55	0.60
Oregon	Connecticut	Iowa	Idaho	Kansas	Massachusetts	North Carolina	Vermont		
	0.71	0.58	0.52	0.95	0.67	0.54	0.78		
Pennsylvania	Alaska	Kentucky	Massachusetts	Montana	Texas				
	1.00	0.55	0.83	0.58	0.64				
Rhode Island	Connecticut	Massachusetts	Maryland	Maine	North Carolina	New Mexico	Nevada		
	0.57	0.68	0.69	0.51	0.85	0.65	0.50		

Response States_ t	Explanatory States_ t								
South Carolina	Arkansas	Idaho	Kansas	Massachusetts	Maryland	Minnesota	Texas	Vermont	
	0.73	0.53	0.53	0.59	0.72	0.62	0.73	0.60	
South Dakota	Colorado	Kentucky	Montana	North Carolina	Nevada	Washington	West Virginia		
	0.56	0.61	0.83	0.56	0.54	0.62	0.66		
Tennessee	Montana	Texas							
	0.53	0.73							
Texas	Vermont								
	0.59								
Utah	Alaska	Arkansas	Iowa	Indiana	Kansas	Kentucky	Montana	New Mexico	Nevada
	0.51	0.50	0.50	0.50	0.68	0.53	0.54	0.57	0.83
Virginia	Connecticut	Massachusetts	Maryland	Maine	New Mexico	Vermont			
	0.80	1.00	0.73	0.80	0.51	0.51			
Vermont	Kansas	Maryland	Minnesota	Tennessee					
	0.50	0.57	0.54	0.55					
Washington	Massachusetts	Maryland	North Carolina	Nevada					
	0.60	0.63	0.74	0.83					
Wisconsin	Alaska	Alabama	Kentucky	Massachusetts	Maryland				
	0.70	0.77	0.99	0.61	0.58				
West Virginia	Alaska	Kentucky	Montana	Ohio	Pennsylvania				
	0.99	0.87	0.50	0.52	0.56				

Bold entries represent the selected edges for the MAR structure based on posterior probabilities greater than 0.5. The states highlighted indicate strong evidence of directed edges originating from explanatory to response states. There are at least nine cases that explanatory states have strong evidence of directed edges to response states.

Table 3 Number of cases that US states are the origination of directed edges (the MAR structure)

Explanatory States	# of being the origination with $p>0.5, h=1$	# of being the origination with $p>0.5, h=2$	# of being the origination with $p>0.5, h=3$	# of being the origination with $p>0.5, h=4$
Alaska	5	8	2	1
Alabama	8	4	9	9
Arkansas	2	3	6	8
Arizona	9	2	2	2
California	9	15	16	16
Colorado	4	8	2	3
Connecticut	5	4	6	4
District of Columbia	8	13	13	13
Delaware	2	9	11	13
Florida	1	5	6	7
Georgia	2	5	3	1
Hawaii	3	5	8	11
Iowa	11	9	14	13
Idaho	13	4	2	5
Illinois	9	7	6	3
Indiana	1	6	4	8
Kansas	15	19	21	28
Kentucky	3	2	2	3
Louisiana	10	14	14	11
Massachusetts	5	6	1	1
Maryland	2	2	1	6
Maine	8	2	3	1
Michigan	2	11	17	14
Minnesota	1	1	2	1
Missouri	9	15	16	14
Mississippi	6	2	3	3
Montana	3	3	7	4
North Carolina	13	12	4	3
North Dacota	4	9	8	7
Nebraska	4	3	3	3
New Hampshire	15	10	12	14
New Jersey	4	3	1	2
New Mexico	6	3	3	2
Nevada	6	2	0	0
New York	11	7	10	9
Ohio	12	3	1	3
Oklahoma	6	5	3	2
Oregon	4	7	12	4
Pennsylvania	3	2	1	1
Rhode Island	8	5	1	0
South Carolina	4	4	3	3
South Dakota	9	0	0	3
Tennessee	1	1	1	1

Explanatory States	# of being the origination with $p>0.5, h=1$	# of being the origination with $p>0.5, h=2$	# of being the origination with $p>0.5, h=3$	# of being the origination with $p>0.5, h=4$
Texas	1	2	3	3
Utah	7	3	4	4
Virginia	3	1	2	4
Vermont	2	1	1	1
Washington	8	3	1	1
Wisconsin	4	7	8	8
West Virginia	4	4	2	2

Table 4 Number of cases that US states are the origination of directed edges (the MIN structure)

Explanatory States	# of being the origination with $p>0.5, h=1$	# of being the origination with $p>0.5, h=2$	# of being the origination with $p>0.5, h=3$	# of being the origination with $p>0.5, h=4$
Alaska	15	1	17	0
Alabama	5	2	0	2
Arkansas	5	0	0	0
Arizona	0	0	0	0
California	1	0	0	3
Colorado	7	3	2	2
Connecticut	5	3	0	0
District of Columbia	0	0	4	1
Delaware	0	1	4	0
Florida	0	3	3	6
Georgia	1	1	2	7
Hawaii	0	3	10	5
Iowa	3	1	3	1
Idaho	12	0	1	17
Illinois	7	4	0	12
Indiana	6	0	0	3
Kansas	13	12	9	5
Kentucky	15	1	0	0
Louisiana	1	5	3	6
Massachusetts	18	3	9	1
Maryland	16	0	0	3
Maine	8	17	14	13
Michigan	1	4	1	0
Minnesota	8	15	0	0
Missouri	0	0	6	8
Mississippi	7	18	0	0
Montana	16	22	8	3
North Carolina	11	11	0	0
North Dakota	0	0	2	1
Nebraska	1	0	6	3
New Hampshire	0	17	15	9
New Jersey	0	6	6	0
New Mexico	7	0	4	0
Nevada	11	15	8	1
New York	1	1	3	5
Ohio	4	2	15	20
Oklahoma	0	3	1	0
Oregon	0	1	2	0
Pennsylvania	1	8	8	8
Rhode Island	2	8	0	0
South Carolina	1	0	0	15
South Dakota	0	24	17	15
Tennessee	6	0	25	4

Explanatory States	# of being the origination with $p>0.5, h=1$	# of being the origination with $p>0.5, h=2$	# of being the origination with $p>0.5, h=3$	# of being the origination with $p>0.5, h=4$
Texas	11	2	2	5
Utah	0	1	0	3
Virginia	0	2	0	0
Vermont	14	0	0	0
Washington	4	8	5	0
Wisconsin	2	5	4	8
West Virginia	3	0	0	0