

Who falls for fake news? Psychological and clinical profiling evidence of fake news consumers

Running title

Who falls for fake news?

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Abstract

Fake news increased after the coronavirus pandemic and its impact on psychopathology and individual differences is still unclear. This research analysed the psychological and psychopathological profiles that characterize the fake news consumer. A total of 1,452 volunteers from the general population with no previous psychiatric history participated. They responded to clinical assessment tests on psychopathology. They had to solve a fake news screening test to be classified into a quasi-experimental group: 1 (non-fake news consumers) or group 2 (fake news consumers). Mean comparison, Bayesian inference and multiple regression analyses were applied. Participants with a schizotypal, paranoid, and histrionic personality were ineffective in detecting fake news and were more vulnerable to suffer its negative effects. Those participants who failed to detect fake news had higher levels of anxiety and committed more cognitive biases based on suggestibility and the Barnum Effect. No significant effects on psychotic symptomatology or affective mood states were observed. According to the results obtained, two clinical and therapeutic recommendations related to the reduction of the Barnum Effect and the reinterpretation of digital media sensationalism were made. The impact of fake news and possible ways of prevention are also discussed.

Keywords: Fake news; Pseudoscientific information; Cognitive biases; Individual differences; Clinical prevention.

1. Introduction

Pseudoscientific beliefs have many definitions (e.g., [Lilienfeld et al., 2005](#); [Matute et al., 2011](#); [Fasce & Adrián-Ventura, 2020](#)). Current research posits, pseudoscientific beliefs are considered to happen if certain content or information is accepted as scientific when in fact it lacks sufficient contrasted evidence ([Fasce & Picó, 2019](#)). Several studies have found significant increases in pseudoscientific beliefs and psychopathological risks during the coronavirus pandemic ([Escolà-Gascón et al., 2020](#)). Pseudoscientific beliefs are frequently irrational (e.g., believing in the existence of the Loch Ness monster), but they do not necessarily have to be magical (e.g., believing in the existence of intelligent extra-terrestrial life) (see [Dagnall et al., 2010a](#); [Dagnall et al., 2010b](#); [Dagnall et al., 2011](#); [Vickers, 2020](#); [Escolà-Gascon et al., 2021a](#)). In this last example, the content is not magical because in the field of astrophysics scientists are looking for and finding a multitude of exoplanets (see the publication of [Samland et al., 2021](#)). However, at present there is no evidence of intelligent life outside the Earth (for this reason the statement "there is extra-terrestrial intelligent life" can be considered pseudoscientific). In this paper, the authors focus on pseudoscientific beliefs related to medicine.

Pseudoscientific beliefs related to health put people's safety and well-being at risk ([Lilienfeld, 2007](#); [Lazarević et al., 2021](#)). As an example, some studies observed that denial of the existence of the AIDS virus encouraged many people to practice sex without the use of protection ([Bogart et al., 2010](#); [Ojikutu et al., 2020](#)). Similarly, other research analysed the importance of vaccination and the possible repercussions of dissident movements with vaccines ([Kata, 2010](#); [Pullan & Dey, 2021](#)). In addition, believing that vaccines are detrimental to health could be considered a pseudoscientific belief ([Jolley & Douglas, 2014](#); [Stein et al., 2021](#)). Moreover, certain studies found that pseudoscientific beliefs are related to conspiracy ideation ([Lobato et al., 2014](#); [Drinkwater et al., 2021](#)).

Conspiracy ideation consists of advocating and believing that alternative theories are real ([Denovan et al., 2020](#)). A conspiracy theory is an implausible argument that offers an unlikely explanation for certain events or circumstances (see [Dagnall et al., 2017a](#); [Drinkwater et al., 2020](#)). For example, believing that governments included a micro-chip in vaccines to control the population (e.g., [Kanozia & Arya, 2021](#)). In clinical terms, conspiracy theories and beliefs are characterized by generating paranoid responses in people ([Darwin et al., 2011](#); [Dagnall et al., 2015](#); [Imhoff & Lamberty, 2018](#)). However, it is unclear what the psychopathological

impact of these paranoid responses are (e.g., [Kuhn et al., 2021](#)). As with the perceptual alternations that characterize magical ideation, it is possible that in conspiracy ideation these levels of paranoia are also manifested both pathologically and nonpathologically ([Kowalski et al., 2020](#)). Therefore, pseudoscientific beliefs may include or be related to conspiracy theories because both constructs lack scientific evidence to ensure their certainty or plausibility ([Escolà-Gascón et al., 2021b](#)).

The concept of "fake news" is defined as content that is disseminated or released as "real" information when it is not ([Lazer et al., 2018; van der Linden et al., 2020](#)). In this sense, fake news is invented (see [Escolà-Gascón et al., 2021](#)). This definition bears much resemblance to the definitions highlighted above in relation to pseudoscientific beliefs and conspiracy theories. Moreover, fake news also lacks evidence ([Lindeman & Aarnio, 2007; Bronstein et al., 2018](#)). In fact, a conspiracy idea could manifest itself in the form of fake news if it became a publicly disseminated piece of information (e.g., conspiracy denial of climate change) (e.g., [van der Linden, 2015](#)). Fake news is problematic because it can socially and psychologically influence populations ([Zimmermann & Kohring, 2020](#)). These changes are difficult to predict because the most frequent media in which fake news are spread is the Internet (see [Guess et al., 2019](#)). Actually, the Internet makes it difficult to discriminate between verified and false information (e.g., [Pennycook & Rand, 2019](#)).

Researchers have undertaken extensive research examining the influence of fake news literature (see [Pennycook & Rand, 2021](#) for a review). One effect of fake news is radicalization of political ideologies ([Vegetti & Mancosu, 2020; Salvi et al., 2021](#)). Polarization of ways of thinking generates an increase in aggressiveness and fosters ideological tribalism (see [Clark & Winegard, 2020](#)). Another psychological consequence is emotional deregulation (e.g., [Preston et al., 2021](#)). [Martel et al. \(2018\)](#) observed that people who tended to believe more in fake news also had high levels of emotional dependence. Similarly, [Pennycook et al. \(2020\)](#) observed that fake news was also related to thinking styles; levels of analytical style were lower in those who believed in fake news (see also [Swami et al., 2013](#)). However, there is very little evidence on the psychiatric consequences or mental disorders fostered by fake news (see [Pennycook & Rand, 2021](#) "Outstanding Questions").

In the medical field, studying the impact of online fake news on mental health is essential in order to develop prevention programs that protect the psychological well-being of the population. It is true that there are several preventive initiatives based on computational methods, warning signals and mathematical algorithms that restrict problematic online content (e.g., [Apuke et al., 2021](#)). However, although these partially protect the population from access to fake news, they do not address the question of how to safeguard the psychological health of mass Internet consumers ([van der Linden et al., 2021](#)). To protect the mental health of populations that accidentally consume fake news, it is important to know in which categories of psychopathology symptoms increase (see [De Coninck et al., 2021](#); [So et al., 2022](#)).

In this research the authors analysed the symptomatic differences observed between two types of groups: (1) people who effectively discriminated online fake news and (2) people who ineffectively discriminated online fake news. The symptomatic differences were based on four personality disorders, symptoms that characterize *psychotic spectrum disorders* (PSD), anxiety and addiction disorders. The main objective of the study was to find out whether symptom levels of the disorders analysed increased or decreased as a function of the individual's ability to detect fake news online.

In accordance with the evidence initially cited and related to the impact of pseudoscientific beliefs on the mental well-being of the population during the coronavirus pandemic, the following hypothesis was put forward: individuals who are better at detecting online fake news obtain lower levels of symptoms of the disorders evaluated with respect to subjects who do not know how to correctly detect fake news.

2. Methods

2.1. Participants

A total of 1,452 volunteers (49% female and 51% male) over 21 years of age (mean age = 29.15; standard deviation = 6.128) participated. Volunteers were asked how many hours per day they used digital devices to be connected to the Internet (computers, smartphones, tablets, and TVs were included). The number of hours of Internet use per day ranged from 6 to 14 (mean = 10.06; standard deviation = 2.573). This question was asked because the number of online fake news a person can receive varies according to the number of hours of Internet connection. Therefore, this was a control question. Similarly, the participants declared that they had no diagnosed psychiatric history and that they did not have any health impairment to answer the surveys.

All volunteers participated freely and did not receive any compensation for partaking in the study. Participation was voluntary and all had to sign a written informed consent before answering the questionnaires. This consent explained all the information about the study and its objectives.

2.2. Materials

2.2.1. State-Trait Anxiety Inventory (STAI)

This scale was initially developed by [Spielberger et al. \(1970\)](#) to evaluate clinical symptoms of anxiety and stress. The STAI consists of 40 items distributed in two dimensions: (1) State-type anxiety (refers to anxiety motivated by the circumstances experienced by the person) and (2) Trait-type anxiety (refers to anxiety motivated by the individual's way of being or personality). To answer the items, the perceived frequency of symptoms must be indicated on a scale ranging from 0 ("none") to 3 ("very much"). The reliability and validity of this scale is excellent and has been replicated in numerous investigations (e.g., [Wiglusz et al., 2019](#)). In this research, the Spanish version published by TEA Ediciones, S.A.U. was used. ([Spielberger et al., 2015](#)). The reliability indices applied to this sample were acceptable ($\alpha > 0.7$ and $\Omega > 0.7$).

2.2.2. Positive and Negative Affect Schedule (PANAS)

This instrument was originally developed by [Watson et al. \(1988\)](#) to measuring affective symptoms. The PANAS has 20 items that are classified into two dimensions: (1) Positive affect (assesses positive emotions and feelings such as enthusiasm, satisfaction, and vitality) and (2) Negative affect (examines negative emotions such as guilt, hostility, sadness, pessimism, and dissatisfaction). For each item, the participant indicates the degree to which he/she experiences each of the contents on a Likert-type scale from 1 ("very little or not at all") to 5 ("extremely"). The PANAS has satisfactory validity and reliability, both in its original version and in subsequent revisions (see [Díaz-García et al., 2020](#)). In this research, the Spanish adaptation of [Ortuño-Sierra et al. \(2015\)](#) was used. The reliability indices used in this sample were good for both dimensions (alpha >0.8 and omega >0.8).

2.2.3. *Multivariable Multiaxial Suggestibility Inventory (MMSI-2)*

The MMSI-2 was developed and published by [Escolà-Gascón \(2020a\)](#) after 8 years of psychometric analysis. This test is broad-spectrum, has 174 items and measures 16 primary clinical variables and 4-5 secondary variables. These scales collect several clinical symptoms based on The *Diagnostic and Statistical Manual of Mental Disorders* (DSM-5) (see [American Psychiatric Association, 2013](#)). The clinical dimensions of the MMSI-2 used in this research were as follows: Inconsistencies (K), Lies (L), Fraud (F), Simulation (Si), Neurasthenia (Nt), Substance Use (Cs), Suggestibility (Su), Thrill-Seeking (Be), Histrionism (Hi), Schizotypy (Ez), Paranoia (Pa), Narcissism (Na), Anomalous Visual/Auditory Phenomena (Pva), Anomalous Tactile Phenomena (Pt), Anomalous Olfactory Phenomena (Po) and Anomalous Cenesthetic Phenomena (Pc). Participants must indicate the degree of agreement on each of the items following the Likert model of 5 response alternatives (1= "strongly disagree" and 5= "strongly disagree"). The MMSI-2 has excellent construct validity (see [Escolà-Gascon, 2020b](#)) and reliability indices greater than 0.9 (using Cronbach's alpha and McDonald's Omega coefficients) ([Escolà-Gascón, 2020a](#)). In this sample, the reliability indices were also excellent for all dimensions (alpha >0.8 and omega >0.8).

2.3. Procedures

The design of this research was correlational and based on group comparison. The recruitment process of the participants was carried out online using social networks:

Facebook, Twitter, LinkedIn, and WhatsApp groups. The duration of the sample collection was 7 months (from March to September 2021). Prior to the survey, each volunteer was presented with an informed consent form that they had to sign (by clicking on the acceptance box on the online form). All materials were administered in Spanish and all subjects stated that they resided in Spain. There were no setbacks or unforeseen events during data collection.

As responses were obtained, data were automatically recorded in a downloadable Microsoft Excel matrix. During the 7 months, the variables sex, age, and number of hours of Internet connection were monitored. When the number of men and women was similar and the summer vacation was over, data was no longer collected.

No missing values were obtained because no question could be left unanswered. When all the responses were collected, the scores for each of the scales and variables of this research were calculated and the results were analyzed.

2.4. Statistical analysis

Statistical analyses were performed using the R programming language, *RStudio* and *JASP* (see [The Jamovi Project, 2021](#)). Multiple regression analysis was applied (with the *residual least squares* method) and Student's t-test was used. Mann-Whitney U tests were also employed, and Bayes Factors (hereafter BF_{10}) were estimated. The a priori probabilities were adjusted to 50% (i.e., they were *equiprobable*). Using simple transformations, the posteriori distribution or $P(H_1/D)$ was estimated from the BF_{10} . These transformations were as follows:

$$BF_{10} = \frac{\int_{\Theta_{H_1}} P(D | \theta_{H_1}, H_1) \cdot \pi(\theta_{H_1} | H_1) d\theta_{H_1}}{\int_{\Theta_{H_0}} P(D | \theta_{H_0}, H_0) \cdot \pi(\theta_{H_0} | H_0) d\theta_{H_0}} \propto \frac{BF_{10}}{BF_{10} + 1} = P(D | H_1)$$

The parameters of the above equation that allows estimation of the BF_{10} are obtained by integration procedures and the result can be interpreted in the odds metric. For this reason, this transformation could be used. BF_{10} equal to or greater than 10 points indicates evidence in favour of the alternative hypothesis of the study.

3. Results

3.1. Mean comparison analysis

The variables were statistically described according to two groups: (1) participants who effectively detected fake news about coronavirus and (2) participants who ineffectively detected fake news. Participants were classified into these groups according to the number of hits obtained in the COVID-19 fake news test (hits ranged from 0 to 18). When the participant obtained a value equal to or higher than 9 points, he/she was classified in group 1 ($n = 699$). When the score was below 9, the participant was classified in group 2 ($n = 753$). This cut-off point was defined according to the median observed in the fake news test (total mean = 9.81; total standard deviation = 4.308; total median = 9). Descriptive statistics are summarized in Table 1.

Table 1. Descriptive statistics for the scales differentiated according to the effectiveness in the detection of fake news.

Clinical variables	Ineffective fake news detection (hits<9) N = 753		Effective fake news detection (hits≥9) N = 699	
	Direct means	Standard deviation	Direct means	Standard deviation
Number of hits in the detection fake news exam	6.07	2.064	13.82	1.671
Trait anxiety (STAI-T)	22.12 (3.69)	3.875 (0.646)	16.69 (2.78)	3.709 (0.618)
State anxiety (STAI-S)	20.21 (3.37)	3.844 (0.641)	13.71 (2.28)	2.935 (0.489)
Positive affect	31.48 (6.62)	5.164 (1.291)	32.22 (6.81)	5.983 (1.496)
Negative affect	17.67 (3.17)	3.373 (0.843)	15.56 (2.64)	3.418 (0.855)
Histrionism (Hi)	36.22 (6)	2.923 (0.562)	31.43 (5.08)	3.194 (0.614)
Schizotypy (Ez)	38.39 (7.59)	3.821 (0.868)	30.5 (5.80)	3.81 (0.866)
Paranoia (Pa)	35.94 (7.73)	3.328 (0.832)	32.17 (6.79)	2.617 (0.654)
Narcissism (Na)	35.16 (6.28)	3.699 (0.771)	33.44 (5.93)	6.036 (1.257)

Visual and auditory Disturbances (Pva)	17.66 (2.88)	2.565 (0.583)	17.73 (2.89)	2.66 (0.605)
Touch disturbances (Pt)	12.93 (2.83)	1.622 (0.579)	12.72 (2.76)	1.692 (0.604)
Olfactory disturbances (Po)	14.78 (3.49)	2.093 (0.747)	14.68 (3.46)	2.041 (0.729)
Cenesthetic disturbances (Pc)	16.09 (3.08)	2 (0.556)	16.41 (3.17)	1.892 (0.525)
Cognitive biases based on the Barnum Effect (Si).	22.53 (7.30)	2.146 (0.894)	15.75 (4.48)	4.267 (1.778)
Suggestibility (Su)	22.19 (6.14)	4.474 (1.598)	18.43 (4.80)	2.864 (1.023)
Search for emotions (Be)	14.07 (5.67)	2.127 (1.329)	12.73 (4.83)	2.3 (1.438)
Substance use and addiction risks (Cs)	13.54 (3.05)	2.174 (0.777)	13.52 (3.04)	2.061 (0.736)

Note: Means and standard deviations on a scale of 0 to 10 are given in parentheses.

Direct scores enable the comparison between the means of group 1 and 2. However, if it is desired to check which clinical variables scored higher than others, the direct scores should be transformed so that all the scales are on the same metric. In this case, given that not all the scales had Spanish normative groups, it was decided to apply a transformation of the scores into a range from 0 to 10.

The following equation was used to modify the direct scores:

$$TS_{10} = \frac{DS_{ij}}{K_{\max_j} \cdot N_j} \cdot 10$$

Where:

TS_{10} = transformed scores over 10;

DS_{ij} = direct scores of the "i" participant and "j" scale.

K_{\max_j} = maximum response alternative in scale "j".

N_j = number of items in scale "j".

When the minimum response alternative on a scale "j" K_{\min_j} is not equal to 0, the following equation should be used:

$$TS(K_{\min_j} \neq 0)_{10} = \frac{DS_{ij} - K_{\max_j}}{(K_{\max_j} - 1) \cdot N_j} \cdot 10$$

The latter transformation was used for the PANAS and MMSI-2 scales. From these formulas, all the scores were within a range of values between 0 and 10. By having the same metric, the mean graph could be made to compare the scores between the clinical scales and to know which of them scored higher. Comparisons between group 1 and group 2 means were made with direct scores. Table 2 shows the results of these comparisons.

Table 2. Means comparisons between participants who effectively detect fake news (>9) and those who do not (<9).

Clinical variables	Welch's t-test	Mann-Whitney U test	Cohen's d ^a	BF ₁₀ $P(H_1 D)$	% Error associated to the BF ₁₀	Measurement of contrast power (1-β) ^b
Trait anxiety (STAI-T)	27.305*	442,951*	1.432	135.483 ~1	1.827e -131	φ= 27.264 (1-β)= 0.99
State anxiety (STAI-S)	36.412*	479,475.5*	1.902	208.402 ~1	4.725e -203	φ= 36.213 (1-β)= 0.99
Positive affect	-2.504	246,348.5	0.132	1.350 0.574	0.008	-
Negative affect	11.826*	350,005.5*	0.621	37.025 ~1	3.069e -30	φ= 11.823 (1-β)= 0.99
Histrionism (Hi)	29.747*	454,763*	1.564	169.237 ~1	4.572e -152	φ= 29.778 (1-β)= 0.99
Schizotypy (Ez)	39.379*	486,687*	2.067	248.281 ~1	6.482e -230	φ= 39.354 (1-β)= 0.99
Paranoia (Pa)	24.074*	432,044*	1.258	111.275 ~1	2.066e -105	φ= 23.952 (1-β)= 0.99
Narcissism (Na)	6.468*	326,555*	0.342	30.877 ~1	1.268e -10	φ= 6.511 (1-β)= 0.99

Visual-auditory Disturbances (Pva)	-0.548	260,369.5	0.029	0.068	0.152	-
Touch disturbances (Pt)	2.412	285,896	0.127	1.054	0.010	-
Olfactory disturbances (Po)	0.947	270,202	0.050	0.092	0.114	-
Cenesthetic disturbances (Pc)	-3.189	243,770.5	0.167	8.709	0.001	-
Cognitive biases based on the Barnum Effect (Si).	37.761*	476,957*	2.004	230.433 ~1	2.296e -223	φ= 38.155 (1-β)= 0.99
Suggestibility (Su)	19.193*	401,409*	1.000	79.534 ~1	2.083e -70	φ= 19.039 (1-β)= 0.99
Search for emotions (Be)	11.469*	360,308.5*	0.603	29.047 ~1	1.012e -28	φ= 11.481 (1-β)= 0.99
Substance use and addiction risks (Cs)	0.130	264,024	0.007	0.059	0.175	-

Note:*p<0.001

a. Cohen's d test was corrected by applying Hedges' g.

b. Since the sample size of the two groups was different, the harmonic mean was used as a parameter to estimate the contrast power of each test. The harmonic mean was 724.996.

The harmonic mean equation is as follows:

$$\bar{n} = \frac{(2n_1n_2)}{(n_1 + n_2)}$$

Mean transformed scores out of 10 were graphically depicted in Figure 1 for each type of group.

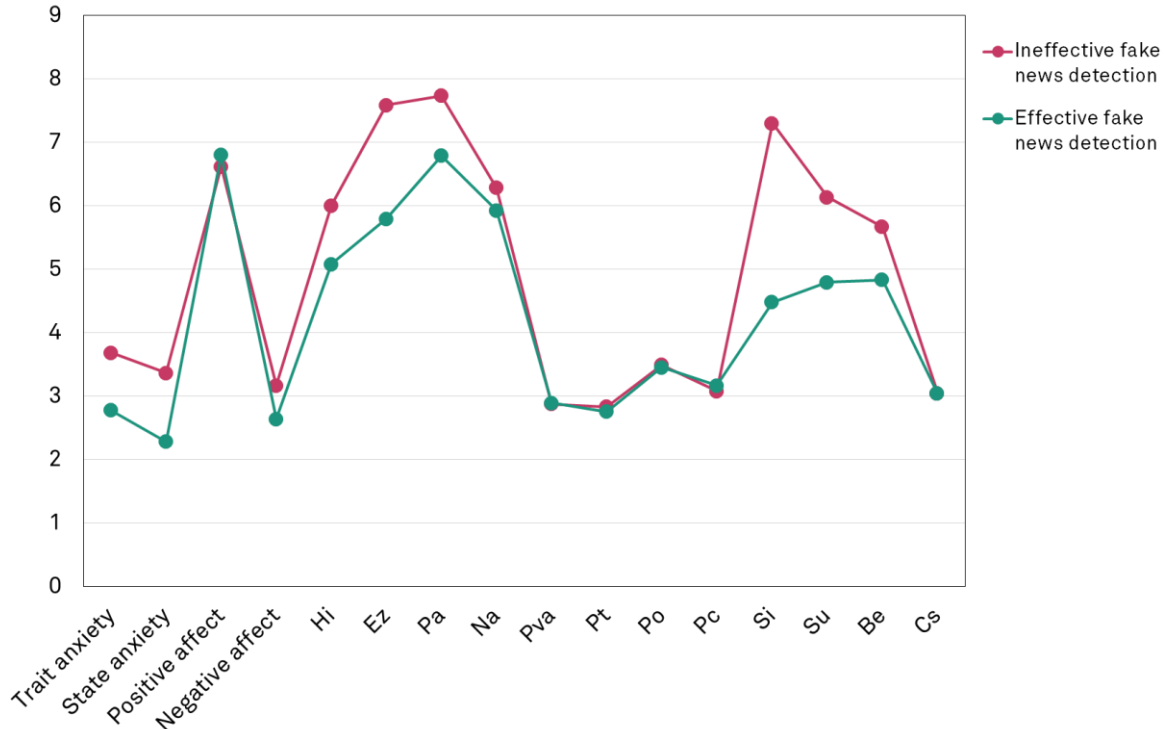


Figure 1. Individual/group differences from means out of 10 for clinical personality tendencies, differentiating between participants who effectively detect fake news and those who do not.

Results in Table 2 indicate that subjects with difficulties in detecting fake news score higher on the subclinical variables Trait Anxiety, State Anxiety, Negative Affect, Histrionism, Schizotypy, Paranoia, Narcissism, Simulations (Barnum Effect), Suggestibility and Search for emotions. The Bayes Factor exceeded the 10-point threshold and reinforced the observations obtained in the frequentist contrasts. The contrast power was excellent for all the statistical tests applied to these variables. This is important because it prevents type I and type II errors.

The graphical trends in Figure 1 can be interpreted as a group clinical profile. The clinical variables that scored highest in group 1 subjects were Positive Affect and Paranoia. In contrast, the variables that scored highest in group 2 were Schizotypy, Simulations and Paranoia. These clinical trends reinforce what was observed in Table 2 and allow us to conclude that not knowing how to detect fake news effectively is related to increases in some clinical variables such as Anxiety, Negative Affect, Simulations (Barnum Effect bias) and personality tendencies.

3.2. Multiple regression analysis

To re-examine the observed statistical effects and trends, a multiple regression analysis was applied using fake news detection hits as the criterion variable and subclinical variables as predictors. Table 3 shows the correlations between fake news test scores and subclinical personality variables, perceptual disturbances and cognitive biases. Based on Table 3, predictors that best correlated with the criterion variable were chosen to fit the regression models. Table 4 provides the results of the forward stepwise regression.

Table 3. Correlations between subclinical (predictor) variables and the fake news detection test score. (N = 1,452).

Clinical variables	Pearson correlation	95% Pearson Confidence intervals		Spearman correlation
		Lower	Upper	
Trait anxiety (STAI-T)	-0.578*	-0.611	-0.542	-0.603*
State anxiety (STAI-S)	-0.671*	-0.699	-0.642	-0.710*
Positive affect	0.157	0.107	0.207	0.191
Negative affect	-0.377*	-0.420	-0.332	-0.410*
Histrionism (Hi)	-0.650*	-0.679	-0.619	-0.695*
Schizotypy (Ez)	-0.748*	-0.770	-0.724	-0.763*
Paranoia (Pa)	-0.711*	-0.736	-0.685	-0.738*
Narcissism (Na)	-0.300*	-0.346	-0.252	-0.335*
Visual-auditory Disturbances (Pva)	0.011	-0.040	0.062	-0.023
Touch disturbances (Pt)	-0.103	-0.154	-0.052	-0.134
Olfactory disturbances (Po)	-0.044	-0.095	0.008	-0.058
Cenesthetic disturbances (Pc)	0.054	0.002	0.105	0.012
Cognitive biases based on the Barnum Effect (Si).	-0.757*	-0.778	-0.734	-0.787*
Suggestibility (Su)	-0.486*	-0.524	-0.446	-0.522*
Search for emotions (Be)	-0.457*	-0.496	-0.415	-0.464*

Substance use and addiction risks (Cs)	-0.071	-0.122	-0.020	-0.127
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Note: *p<0.001.

Table 4. Multiple linear regression models using the forward stepwise method. The inclusion of predictor variables was based on linear correlations. Criterion variable = fake news detection test hits; predictor variables: State anxiety, Histrionics, Schizotypy, Paranoia and Simulations (Barnum Effect).

Clinical variables	Unstandardized regression parameters			β_z	Adjusted R ² (error of the estimate)	ΔR^2	Fisher's test change
	β_0	β	Errors				
Model 1							
Cognitive biases based on the Barnum Effect (Si).	23.017	-0.686*	0.016	-0.757*	0.573	-	1,944.374*
Model 2							
Cognitive biases based on the Barnum Effect (Si).	31.136	-0.445*	0.015	-0.491*	0.723	0.150	785.084*
Schizotypy (Ez)		-0.369*	0.013	-0.470*			
Model 3							
Cognitive biases based on the Barnum Effect (Si).	37.396	-0.358*	0.015	-0.395*	0.758	0.036	216.074*
Schizotypy (Ez)		-0.292*	0.013	-0.371*			
Paranoia (Pa)		-0.311*	0.021	-0.256*			
Model 4							
Cognitive biases based on the Barnum Effect (Si).	36.141	-0.324*	0.015	-0.357*	0.775	0.017	107.262*
Schizotypy (Ez)		-0.252*	0.013	-0.321*			
Paranoia (Pa)		-0.255*	0.021	-0.210*			
State anxiety (STAI-S)		-0.157*	0.015	-0.172*			

Model 5							
Cognitive biases based on the Barnum Effect (Si).		-0.318*	0.015	-0.351*			
Schizotypy (Ez)	37.737	-0.218*	0.014	-0.278*	0.781	0.006	40.597*
Paranoia (Pa)		-0.227*	0.021	-0.187*			
State anxiety (STAI-S)		-0.142*	0.015	-0.156*			
Histrionism (Hi)		-0.121*	0.019	-0.109*			

Note: * $p < 0.001$. β_0 = constant; β = unstandardized regression coefficient; β_z = standardized regression coefficient; R^2 = explained variance.

Results in Table 4 indicate that the variables Simulations (Si), Schizotypy (Ez), Paranoia (Pa), State anxiety and Histrionism (Hi) explain 78.1% of the variance of the fake news test scores. However, the increases in R^2 and Fisher's F-statistic indicate that the most important increase can be observed with the inclusion of the variable Schizotypy (Ez) (see model 2). This has several clinical implications that will be analyzed in the discussion.

4. Discussion

The purpose of this research was to investigate the effects of fake news on subclinical personality profiles and other psychopathological variables. Results support the conclusion that failure to correctly detect fake news is related to increased psychopathological risks in the following variables: Trait Anxiety, State Anxiety, Negative Effect, Histrionism, Schizotypy, Paranoia, Narcissism, Simulations (Barnum Effect), Suggestibility and Search for Emotions. The implications of these results for the clinical practice of psychology and psychiatry are discussed and interpreted below.

4.1. Connection between fake news and personality disorders

The relationship between fake news and levels of schizotypy was consistent with previous scientific evidence on pseudoscientific beliefs and magical ideation (see [Bronstein et al., 2019](#); [Escolà-Gascón et al., 2021b](#)). Following the dual process theory model (e.g., [Pennycook & Rand, 2019](#)), when a person does not correctly distinguish between information with scientific arguments and information without scientific grounds it is because he or she

predominantly uses cognitive reasoning characterized by intuition (e.g., [Dagnall et al., 2010](#); [Swami et al., 2013](#); [Dagnall et al., 2017b](#); [Williams et al., 2021](#)).

Concomitantly, use of intuitive thinking was positively correlated with magical beliefs (see [Šrol, 2021](#)). Psychopathological classifications include magical beliefs as a dimension of schizotypal personality (e.g., [Escolà-Gascón, 2020a](#)). Therefore, it is possible that the high schizotypy scores in this study can be explained from the perspective of dual process theory ([Denovan et al., 2018](#); [Denovan et al., 2020](#); [Drinkwater et al. 2021](#)): intuitive thinking could be the moderating variable explaining why participants who scored higher in schizotypy did not effectively detect fake news.

Something similar happened with the subclinical trait of paranoia. This variable scored the highest in both group 1 and group 2 (see Figure 1). Intuition is also positively related to conspiratorial ideation (see [Drinkwater et al., 2020](#); [Gligorić et al., 2021](#)). Similarly, psychopathology tends to classify conspiracy ideation as a frequent belief system in paranoid personality (see [Escolà-Gascón et al., 2022](#)). This is because conspiracy beliefs are based on systematic distrust of the systems that structure society (political system), knowledge (science) and economy (capitalism) ([Swami et al., 2013](#); [Dagnall et al., 2015](#)). Likewise, it is known that distrust is the transversal characteristic of paranoid personality ([So et al., 2022](#)). Then, in this case the use of intuitive thinking and dual process theory could also justify the obtained paranoia scores. The same is not true for the histrionic personality.

One of the characteristic features of histrionicity is lying (see [Escolà-Gascón, 2020a](#)). Specifically, within the histrionic personality, lying should be understood as an automatic deception that is used for seduction purposes ([Posavac et al., 2021](#)). In this context, psychological theories that justify the functionality of unconscious lying could serve us to understand why histrionic scores were higher in the group that ineffectively detected fake news (see [Semrad et al., 2019](#)). One such theory is social desirability ([Edwards, 1957](#); [Danioni et al., 2021](#)). It is possible that participants who did not effectively detect fake news believed that this research supported conspiratorial beliefs. With this prior instruction, this type of participant might wish to please and correspond with the expectations of the research team. This kind of behavior is very common in histrionic profiles (see [Furnham & Grover, 2021](#)). For this reason, histrionic scores were elevated in those participants who did not correctly detect fake news.

STAI scales also showed differences between group 1 and group 2. Participants who did not effectively detect fake news (vs. effectively detected fake news) had higher levels of anxiety than participants. In this relationship, anxiety can be understood as a response to the information and misinformation overload during the coronavirus pandemic (e.g., [Gupta et al., 2022](#)), or as an independent variable that could distort the individual's reasoning for identifying fake news (see [Escolà-Gascón et al., 2020](#)). In this research, anxiety was measured both as a stable personality trait and as a state. The following speculation could be made, it is possible that state-type anxiety levels reflect a circumstantial response to the coronavirus pandemic and that trait-type anxiety assumes the role of an independent variable.

However, this point is only a speculation; it is crucial to note that the above speculation could not be experimentally tested in the present study. For this reason, in the regression models it was accepted that state-type anxiety acted as a predictor variable; it would also be possible that the circumstances of confrontation with fake news (we refer to the fact of having to answer a fake news screening test) increased the levels of situational anxiety and caused an increase in the responses of some participants. Therefore, the interpretation of the value of the anxiety scores should be made with caution.

Narcissistic personality and scales measuring perceptual disturbances did not yield relevant effect sizes. The fact that perceptual disturbances were not significant subclinical variables implies that positive psychotic symptomatology (i.e., those symptoms related to hallucinations and delusions) are not affected by fake news. This would only be applicable for positive symptoms, as negative symptomatology (included in schizotypy) should have been assessed independently (see the limitations section). The same is true for the positive and negative affect scales, which showed no significant differences.

4.2. Clinical recommendations for the prevention and detection of fake news

Based on the findings, prevention of fake news should be carried out both externally (for example, using mathematical algorithms to detect and eliminate fake news) and from an individual or internal approach. The latter approach refers to the internal psychological resources used by the user of social networks to detect fake news and avoid the discomfort they generate. Encouraging the use and learning of these internal resources is also a way of

combating fake news. In this line, the results obtained in this study allow us to make several contributions.

First, the Simulations (Si) scale of the MMSI-2 was one of the variables that showed the greatest differences. This scale is related to the Barnum Effect, which is a type of cognitive bias. The Barnum Effect consists of accepting as exclusive a verbal description of an individual's personality, when, the description employs contents applicable or generalizable to any profile or personality that one wishes to describe (see [Boyce & Geller, 2002](#); [O'Keeffe & Wiseman, 2005](#)). The error of this bias is to assume as exclusive or unique information that is not. This error can occur in other contexts not limited to personality descriptions. Originally, this bias was studied in the field of horoscopes and pseudoscience's (see [Matute et al., 2011](#)). Research results suggest that people who do not effectively detect fake news regularly commit the Barnum Effect. So, one way to prevent fake news may be to educate about what the Barnum Effect is and how to avoid it. The question we ask is this, by reducing the Barnum effect, could the negative impact of fake news be avoided?

According to scores on the Suggestibility and Thrill-Seeking scales, it seems that people who score high are those who engage with higher levels of fake news. This could indicate and describe the vulnerable profile of those who are more susceptible to the negative effects of fake news. Therefore, this proposes two possible recommendations: (1) consistent with [Pennycook et al. \(2018\)](#), educate the interpretation of sensationalism associated with fake news to interfere as little as possible in those users with high Suggestibility, and (2) as suggested by [Bronstein et al. \(2019\)](#), foster critical thinking as an internal psychological resource that questions and allows dismantling fake news. This coincides with the evidence obtained by [Bago et al. \(2020\)](#), which makes our results consistent with the findings of other studies (see also [Roozenbeek & van der Linden, 2019](#)). Moreover, an adaptive use of critical thinking could help suggestible individuals to better regulate their emotional lability (see [Escolà-Gascón et al., 2021c](#)).

The promotion of critical thinking involves enhancing analytical cognitive reasoning, which is complementary to the intuitive style. (see [Pennycook et al., 2020](#)). This idea based on dual process theory supports the interpretations in subsection 4.1. At this point, it is important to remember that the activation of the analytical style does not imply annulling the use of the intuitive style. Similarly, when intuition is used, it is also possible to active the

analytical-critical style simultaneously. In this way, dual processing models suggest that both styles function in parallel ([Pennycook & Rand, 2021](#)). Therefore, the promotion of critical thinking should be based on analytical cognitive reasoning, but without forgetting that intuitive reasoning can intervene at specific moments. This is frequent in individuals who believe in the existence of the paranormal. These profiles tend to use intuition as a decision-making mechanism (see [Drinkwater et al., 2021](#); [Williams et al., 2021](#)), but they can also think critically and activate analytical reasoning. Thus, the use of critical thinking based exclusively on the analytical style should be taken with caution.

These suggestions can be applied at several levels. One of these levels may be the psychotherapeutic and clinical level. Caution should be exercised; accordingly, the above suggestions are presented only as hypothetical ideas that should be tested again in the future to know and verify their efficacy.

4.3. Limitations

One limitation has already been discussed in section 4.1. However, two more can be highlighted. The first is related to the effective and ineffective detection of fake news. It should be considered that the cut-off point chosen to organize groups 1 and 2 was empiricist because it was based on the median. However, there is no previous research indicating the sensitivity and specificity of the chosen threshold. As these statistical properties are unknown, the results should be considered with caution and replicated in future research. In addition, another sub-limitation should also be highlighted: the fact that a person does not effectively detect fake news does not mean that he/she has magical or conspiratorial ideation. This possibility was not tested in this research and should be analyzed in future studies. We hypothesize that failure to detect pseudoscientific fake news will be related to magical ideation and intuitive thinking. However, the ideal would be to contrast this possibility from the logic of experimental designs.

The second limitation is related to the generalization of the results and the sample. Although the sample size was excellent for this type of research, all subjects were Spanish. It is likely that cultural and/or linguistic differences may interfere with the results and moderate the effects of the relationship between fake news detection and subclinical attributes.

4.4. Conclusions

The conclusions of this research can be summarized as follows:

(1) The evidence obtained proposes that profiles with high scores in schizotypy, paranoia and histrionism are more vulnerable to the negative effects of fake news. In clinical practice, special caution is recommended for patients who meet the symptomatic characteristics of these personality traits.

(2) In psychiatry and clinical psychology, it is proposed to combat fake news by reducing or recoding the Barnum effect, reinterpreting sensationalism in the media and promoting critical thinking in social network users. These suggestions can be applied from intervention programs but can also be implemented as psychoeducational programs for massive users of social networks.

(3) Individuals who do not effectively detect fake news tend to have higher levels of anxiety, both state and trait anxiety. These individuals are also highly suggestible and tend to seek strong emotions. Profiles of this type may inappropriately employ intuitive thinking, which could be the psychological mechanism that

(4) Positive psychotic symptomatology, affective mood states and substance use (addiction risks) were not affected by fake news. In the field of psychosis, it should be analyzed whether fake news influences negative psychotic symptomatology.

The results of this research are not experimental, but they contribute to the generation of new hypotheses and offer practical recommendations for the psychiatric and/or psychological clinic.

References

- American Psychiatric Association. (2013). *Diagnostic and Statistical Manual of Mental Disorders (DSM-5)*. (5th ed.). American Psychiatric Association.
- Apuke, O., & Omar, B. (2021). Fake news and COVID-19: modelling the predictors of fake news sharing among social media users. *Telematics and Informatics*, *56*, 101475. <https://doi.org/10.1016/j.tele.2020.101475>
- Bago, B., Rand, D. G., & Pennycook, G. (2020). Fake news, fast and slow: Deliberation reduces belief in false (but not true) news headlines. *Journal of Experimental Psychology: General*, *149*(8), 1608-1613. <https://doi.org/10.1037/xge0000729>
- Clark, C., & Winegard, B. (2020). Tribalism in War and Peace: The Nature and Evolution of Ideological Epistemology and Its Significance for Modern Social Science. *Psychological Inquiry*, *31*(1), 1-22. <https://doi.org/10.1080/1047840x.2020.1721233>
- Bogart, L., Wagner, G., Galvan, F., & Banks, D. (2010). Conspiracy Beliefs About HIV Are Related to Antiretroviral Treatment Nonadherence Among African American Men With HIV. *JAIDS Journal of Acquired Immune Deficiency Syndromes*, *53*(5), 648-655. <https://doi.org/10.1097/qai.0b013e3181c57dbc>
- Boyce, T. E., & Geller, E. S. (2002). Using the Barnum effect to teach psychological research methods. *Teaching of Psychology*, *29*(4), 316-318.
- Bronstein, M., Pennycook, G., Bear, A., Rand, D., & Cannon, T. (2019). Belief in Fake News is Associated with Delusionality, Dogmatism, Religious Fundamentalism, and Reduced Analytic Thinking. *Journal of Applied Research in Memory and Cognition*, *8*(1), 108-117. <https://doi.org/10.1016/j.jarmac.2018.09.005>
- Dagnall, N., Denovan, A., Drinkwater, K., Parker, A., & Clough, P. (2017). Statistical bias and endorsement of conspiracy theories. *Applied Cognitive Psychology*, *31*(4), 368-378. <https://doi.org/10.1002/acp.3331>
- Dagnall, N., Denovan, A., Drinkwater, K., Parker, A., & Clough, P. J. (2017). Urban legends and paranormal beliefs: the role of reality testing and schizotypy. *Frontiers in Psychology*, *8*. Article 942. <https://doi.org/10.3389/fpsyg.2017.00942>
- Dagnall, N., Drinkwater, K., & Parker, A. (2011). Alien visitation, extra-terrestrial life, and paranormal beliefs. *Journal of Scientific Exploration*, *25*(4), 699-720.

- Dagnall, N., Drinkwater, K., Parker, A., Denovan, A., & Parton, M. (2015). Conspiracy theory and cognitive style: a worldview. *Frontiers in Psychology*, *6*. Article 206.
<https://doi.org/10.3389/fpsyg.2015.00206>
- Dagnall, N., Drinkwater, K., Parker, A., & Munley, G. (2010a). Reality testing, belief in the paranormal, and urban legends. *European Journal of Parapsychology*, *25*, 25-55
- Dagnall, N., Munley, G., Parker, A., & Drinkwater, K. (2010b). The relationship between belief in extra-terrestrial life, UFO-related beliefs and paranormal belief. *Journal of the Society for Psychical Research*, *74*(898), 1-14.
- Dagnall, N., Parker, A., Munley, G., & Drinkwater, K. (2010). Common paranormal belief dimensions. *Journal of Scientific Exploration*, *24*(3), 431-477.
- Danioni, F., Coen, S., Rosnati, R., & Barni, D. (2020). The relationship between direct and indirect measures of values: Is social desirability a significant moderator? *European Review of Applied Psychology*, *70*(3), 100524.
<https://doi.org/10.1016/j.erap.2020.100524>
- Darwin, H., Neave, N., & Holmes, J. (2011). Belief in conspiracy theories. The role of paranormal belief, paranoid ideation and schizotypy. *Personality and Individual Differences*, *50*(8), 1289-1293. <https://doi.org/10.1016/j.paid.2011.02.027>
- De Coninck, D., Frissen, T., Matthijs, K., d'Haenens, L., Lits, G., & Champagne-Poirier, O. et al. (2021). Beliefs in Conspiracy Theories and Misinformation About COVID-19: Comparative Perspectives on the Role of Anxiety, Depression and Exposure to and Trust in Information Sources. *Frontiers in Psychology*, *12*.
<https://doi.org/10.3389/fpsyg.2021.646394>
- Denovan, A., Dagnall, N., Drinkwater, K., & Parker, A. (2018). Latent profile analysis of schizotypy and paranormal belief: Associations with probabilistic reasoning performance. *Frontiers in Psychology*, *9*. Article 35.
<https://doi.org/10.3389/fpsyg.2018.00035>
- Denovan, A., Dagnall, N., Drinkwater, K., Parker, A., & Neave, N. (2020). Conspiracist beliefs, intuitive thinking, and schizotypal facets: A further evaluation. *Applied Cognitive Psychology*, *34*(6), 1394-1405. <https://doi.org/10.1002/acp.3716>
- Díaz-García, A., González-Robles, A., Mor, S., Mira, A., Quero, S., & García-Palacios, A. et al. (2020). Positive and Negative Affect Schedule (PANAS): psychometric properties

- of the online Spanish version in a clinical sample with emotional disorders. *BMC Psychiatry*, 20(1). <https://doi.org/10.1186/s12888-020-2472-1>
- Drinkwater, K., Dagnall, N., Denovan, A., & Neave, N. (2020). Psychometric assessment of the Generic Conspiracist Beliefs Scale. *PLoS ONE*, 15(3), e0230365. <https://doi.org/10.1371/journal.pone.0230365>
- Drinkwater, K. G., Dagnall, N., Denovan, A., & Walsh, R. S. (2021). To what extent have conspiracy theories undermined COVID-19: Strategic narratives? *Frontiers in Communication*, 6, 47. <https://doi.org/10.3389/fcomm.2021.576198>
- Drinkwater, K. G., Dagnall, N., Denovan, A., & Williams, C. (2021). Paranormal belief, thinking style and delusion formation: a latent profile analysis of within-individual variations in experience-based paranormal facets. *Frontiers in Psychology*, 12. <https://doi.org/10.3389/fpsyg.2021.670959>
- Edwards, A. L. (1957). *The social desirability variable in personality assessment and research*. Dryden Press.
- Escolà-Gascón, Á. (2020a). Researching unexplained phenomena: empirical-statistical validity and reliability of the Multivariable Multiaxial Suggestibility Inventory-2 (MMSI-2). *Heliyon*, 6(7), e04291. <https://doi.org/10.1016/j.heliyon.2020.e04291>
- Escolà-Gascón, Á. (2020b). Researching unexplained phenomena II: new evidences for anomalous experiences supported by the Multivariable Multiaxial Suggestibility Inventory-2 (MMSI-2). *Current Research in Behavioral Sciences*, 1, 100005. <https://doi.org/10.1016/j.crbeha.2020.100005>
- Escolà-Gascón, Á., Marín, F., Rusiñol, J., & Gallifa, J. (2020). Pseudoscientific beliefs and psychopathological risks increase after COVID-19 social quarantine. *Globalization and Health*, 16(1). <https://doi.org/10.1186/s12992-020-00603-1>
- Escolà-Gascón, Á. (2021). New techniques to measure lie detection using COVID-19 fake news and the Multivariable Multiaxial Suggestibility Inventory-2 (MMSI-2). *Computers in Human Behavior Reports*, 3, 100049. <https://doi.org/10.1016/j.chbr.2020.100049>
- Escolà-Gascón, Á., O'Neill, M., & Gallifa, J. (2021a). Beliefs and opinions about the existence of life outside the earth: The UFO Experiences Questionnaire (UFO-Q). *Social Sciences & Humanities Open*, 3(1), 100124. <https://doi.org/10.1016/j.ssaho.2021.100124>

- Escolà-Gascón, Á., Marín, F., Rusiñol, J., & Gallifa, J. (2021b). Evidence of the psychological effects of pseudoscientific information about COVID-19 on rural and urban populations. *Psychiatry Research*, 295, 113628. <https://doi.org/10.1016/j.psychres.2020.113628>
- Escolà-Gascón, Á., Dagnall, N., & Gallifa, J. (2021). Critical thinking predicts reductions in Spanish physicians' stress levels and promotes fake news detection. *Thinking Skills and Creativity*, 42, 100934. <https://doi.org/10.1016/j.tsc.2021.100934>
- Escolà-Gascón, Á. (2022). Impact of conspiracist ideation and psychotic-like experiences in patients with schizophrenia during the COVID-19 crisis. *Journal of Psychiatric Research*, 146, 135-148. <https://doi.org/10.1016/j.jpsychires.2021.12.022>
- Fasce, A., Adrián-Ventura, J., & Avendaño, D. (2020). Do as the Romans do: On the authoritarian roots of pseudoscience. *Public Understanding of Science*, 29(6), 597-613. <https://doi.org/10.1177/0963662520935078>
- Fasce, A., & Picó, A. (2019). Conceptual foundations and validation of the Pseudoscientific Belief Scale. *Applied Cognitive Psychology*, 33(4), 617-628. <https://doi.org/10.1002/acp.3501>
- Furnham, A., & Grover, S. (2021). Do you have to be mad to believe in conspiracy theories? Personality disorders and conspiracy theories. *International Journal of Social Psychiatry*. Advance online publication. Article 002076402110316. <https://doi.org/10.1177/00207640211031614>
- Gligorić, V., Silva, M., Eker, S., Hoek, N., Nieuwenhuijzen, E., Popova, U., & Zeighami, G. (2021). The usual suspects: How psychological motives and thinking styles predict the endorsement of well-known and COVID -19 conspiracy beliefs. *Applied Cognitive Psychology*, 35(5), 1171-1181. <https://doi.org/10.1002/acp.3844>
- Guess, A., Nagler, J., & Tucker, J. (2019). Less than you think: Prevalence and predictors of fake news dissemination on Facebook. *Science Advances*, 5(1), eaau4586. <https://doi.org/10.1126/sciadv.aau4586>
- Gupta, A., Li, H., Farnoush, A., & Jiang, W. (2022). Understanding patterns of COVID infodemic: A systematic and pragmatic approach to curb fake news. *Journal of Business Research*, 140, 670-683. <https://doi.org/10.1016/j.jbusres.2021.11.032>

- Imhoff, R., & Lamberty, P. (2018). How paranoid are conspiracy believers? Toward a more fine-grained understanding of the connect and disconnect between paranoia and belief in conspiracy theories. *European Journal of Social Psychology*, 48(7), 909-926.
<https://doi.org/10.1002/ejsp.2494>
- Jolley, D., & Douglas, K. (2014). The Effects of Anti-Vaccine Conspiracy Theories on Vaccination Intentions. *Plos ONE*, 9(2), e89177.
<https://doi.org/10.1371/journal.pone.0089177>
- Kanozia, R., & Arya, R. (2021). "Fake news", religion, and COVID-19 vaccine hesitancy in India, Pakistan, and Bangladesh. *Media Asia*. Advance online publication.
<https://doi.org/10.1080/01296612.2021.1921963>
- Kata, A. (2010). A postmodern Pandora's box: Anti-vaccination misinformation on the Internet. *Vaccine*, 28(7), 1709-1716. <https://doi.org/10.1016/j.vaccine.2009.12.022>
- Kowalski, J., Marchlewska, M., Molenda, Z., Górska, P., & Gawęda, Ł. (2020). Adherence to safety and self-isolation guidelines, conspiracy and paranoia-like beliefs during COVID-19 pandemic in Poland - associations and moderators. *Psychiatry Research*, 294, 113540. <https://doi.org/10.1016/j.psychres.2020.113540>
- Kuhn, S., Lieb, R., Freeman, D., Andreou, C., & Zander-Schellenberg, T. (2021). Coronavirus conspiracy beliefs in the German-speaking general population: endorsement rates and links to reasoning biases and paranoia. *Psychological Medicine*. Advance online publication. <https://doi.org/10.1017/s0033291721001124>
- Lazarević, L., Purić, D., Teovanović, P., Lukić, P., Zupan, Z., & Knežević, G. (2021). What drives us to be (ir)responsible for our health during the COVID-19 pandemic? The role of personality, thinking styles, and conspiracy mentality. *Personality and Individual Differences*, 176, 110771. <https://doi.org/10.1016/j.paid.2021.110771>
- Lazer, D., Baum, M., Benkler, Y., Berinsky, A., Greenhill, K., & Menczer, F. et al. (2018). The science of fake news. *Science*, 359(6380), 1094-1096.
<https://doi.org/10.1126/science.aao2998>
- Lilienfeld, S. O. (2007). Psychological Treatments That Cause Harm. *Perspectives on Psychological Science*, 2(1), 53-70. <https://doi.org/10.1111/j.1745-6916.2007.00029.x>
- Lilienfeld, S. O., Fowler, K. A., Lohr, J. M., & Lynn, S. J. (2005). Pseudoscience, Nonscience, and Nonsense in Clinical Psychology: Dangers and Remedies. In R. H.

- Wright & N. A. Cummings (Eds.), *Destructive trends in mental health: The well-intentioned path to harm* (pp. 187-218). Routledge.
- Lindeman, M., & Aarnio, K. (2007). Superstitious, magical, and paranormal beliefs: An integrative model. *Journal of Research in Personality, 41*(4), 731-744.
<https://doi.org/10.1016/j.jrp.2006.06.009>
- Lobato, E., Mendoza, J., Sims, V., & Chin, M. (2014). Examining the Relationship Between Conspiracy Theories, Paranormal Beliefs, and Pseudoscience Acceptance Among a University Population. *Applied Cognitive Psychology, 28*(5), 617-625.
<https://doi.org/10.1002/acp.3042>
- Martel, C., Pennycook, G., & Rand, D. (2020). Reliance on emotion promotes belief in fake news. *Cognitive Research: Principles and Implications, 5*(1).
<https://doi.org/10.1186/s41235-020-00252-3>
- Matute, H., Yarritu, I., & Vadillo, M. (2011). Illusions of causality at the heart of pseudoscience. *British Journal of Psychology, 102*(3), 392-405.
<https://doi.org/10.1348/000712610x532210>
- Ojikutu, B., Amutah-Onukagha, N., Mahoney, T., Tibbitt, C., Dale, S., Mayer, K., & Bogart, L. (2020). HIV-Related Mistrust (or HIV Conspiracy Theories) and Willingness to Use PrEP Among Black Women in the United States. *AIDS and Behavior, 24*(10), 2927-2934. <https://doi.org/10.1007/s10461-020-02843-z>
- O’Keeffe, C., & Wiseman, R. (2005). Testing alleged mediumship: Methods and results. *British Journal of Psychology, 96*, 165-179. <https://doi.org/10.1348/000712605X36361>
- Ortuño-Sierra, J., Santarén-Rosell, M., de Albéniz, A., & Fonseca-Pedrero, E. (2015). Dimensional structure of the Spanish version of the Positive and Negative Affect Schedule (PANAS) in adolescents and young adults. *Psychological Assessment, 27*(3), e1-e9. <https://doi.org/10.1037/pas0000107>
- Pennycook, G., Cannon, T. D., & Rand, D. G. (2018). Prior exposure increases perceived accuracy of fake news. *Journal of Experimental Psychology: General, 147*(12), 1865-1880. <https://doi.org/10.1037/xge0000465>
- Pennycook, G., McPhetres, J., Zhang, Y., Lu, J., & Rand, D. (2020). Fighting COVID-19 Misinformation on Social Media: Experimental Evidence for a Scalable Accuracy-

- Nudge Intervention. *Psychological Science*, 31(7), 770-780.
<https://doi.org/10.1177/0956797620939054>
- Pennycook, G., & Rand, D. (2019). Who falls for fake news? The roles of bullshit receptivity, overclaiming, familiarity, and analytic thinking. *Journal of Personality*, 88(2), 185-200. <https://doi.org/10.1111/jopy.12476>
- Pennycook, G., & Rand, D. (2021). The Psychology of Fake News. *Trends in Cognitive Sciences*, 25(5), 388-402. <https://doi.org/10.1016/j.tics.2021.02.007>
- Posavac, S., Kardes, F., Posavac, H., & Gaffney, D. (2021). The Utility of Clinical Psychology Concepts for Judgment and Decision-Making Research: The Case of Histrionic Features. *Personality and Social Psychology Bulletin*, 48(1), 65-77.
<https://doi.org/10.1177/0146167220980887>
- Preston, S., Anderson, A., Robertson, D., Shephard, M., & Huhe, N. (2021). Detecting fake news on Facebook: The role of emotional intelligence. *PLoS ONE*, 16(3). Article e0246757. <https://doi.org/10.1371/journal.pone.0246757>
- Pullan, S., & Dey, M. (2021). Vaccine hesitancy and anti-vaccination in the time of COVID-19: A Google Trends analysis. *Vaccine*, 39(14), 1877-1881.
<https://doi.org/10.1016/j.vaccine.2021.03.019>
- Roozenbeek, J., & van der Linden, S. (2019). Fake news game confers psychological resistance against online misinformation. *Palgrave Communications*, 5(1).
<https://doi.org/10.1057/s41599-019-0279-9>
- Salvi, C., Iannello, P., Cancer, A., McClay, M., Rago, S., Dunsmoor, J., & Antonietti, A. (2021). Going Viral: How Fear, Socio-Cognitive Polarization and Problem-Solving Influence Fake News Detection and Proliferation During COVID-19 Pandemic. *Frontiers in Communication*, 5. <https://doi.org/10.3389/fcomm.2020.562588>
- Samland, M., Bouwman, J., Hogg, D., Brandner, W., Henning, T., & Janson, M. (2021). TRAP: a temporal systematics model for improved direct detection of exoplanets at small angular separations. *Astronomy & Astrophysics*, 646. Article A24.
<https://doi.org/10.1051/0004-6361/201937308>
- Semrad, M., Scott-Parker, B., & Nagel, M. (2019). Personality traits of a good liar: A systematic review of the literature. *Personality and Individual Differences*, 147, 306-316. <https://doi.org/10.1016/j.paid.2019.05.007>

- So, S., Zhu, C., Lincoln, T., Gaudiano, B., Kingston, J., Ellett, L., & Morris, E. (2022). Pandemic paranoia, general paranoia, and their relationships with worry and beliefs about self/others – A multi-site latent class analysis. *Schizophrenia Research*, 241, 122-129. <https://doi.org/10.1016/j.schres.2022.01.045>
- Spielberger, C., Gorsuch, R., & Lushene, R. (1970). *Manual for the State-Trait Anxiety Inventory*. Consulting Psychologists Press.
- Spielberger, C., Gorsuch, R., Lushene, R. (2015). *Cuestionario de Ansiedad Estado-Rasgo, STAI (manual)*. TEA Ediciones, S.A.U. <http://www.web.teaediciones.com/ejemplos/stai-manual-extracto.pdf>
- Šrol, J. (2021). Individual differences in epistemically suspect beliefs: the role of analytic thinking and susceptibility to cognitive biases. *Thinking & Reasoning*, 28(1), 125-162. <https://doi.org/10.1080/13546783.2021.1938220>
- Stein, R., Ometa, O., Pachtman Shetty, S., Katz, A., Popitiu, M., & Brotherton, R. (2021). Conspiracy theories in the era of COVID-19: A tale of two pandemics. *International Journal of Clinical Practice*, 75(2). <https://doi.org/10.1111/ijcp.13778>
- Swami, V., Voracek, M., Stieger, S., Tran, U., & Furnham, A. (2014). Analytic thinking reduces belief in conspiracy theories. *Cognition*, 133(3), 572-585. <https://doi.org/10.1016/j.cognition.2014.08.006>
- van der Linden, S. (2015). The conspiracy-effect: Exposure to conspiracy theories (about global warming) decreases pro-social behavior and science acceptance. *Personality and Individual Differences*, 87, 171-173. <https://doi.org/10.1016/j.paid.2015.07.045>
- van der Linden, S., Roozenbeek, J., & Compton, J. (2020). Inoculating Against Fake News About COVID-19. *Frontiers in Psychology*, 11. <https://doi.org/10.3389/fpsyg.2020.566790>
- van der Linden, S., Roozenbeek, J., Maertens, R., Basol, M., Kácha, O., Rathje, S., & Traberg, C. (2021). How Can Psychological Science Help Counter the Spread of Fake News? *The Spanish Journal of Psychology*, 24. <https://doi.org/10.1017/sjp.2021.23>
- Vegetti, F., & Mancosu, M. (2020). The Impact of Political Sophistication and Motivated Reasoning on Misinformation. *Political Communication*, 37(5), 678-695. <https://doi.org/10.1080/10584609.2020.1744778>

- Vickers, P. (2020). Expecting the unexpected in the search for extraterrestrial life. *International Journal of Astrobiology, 19*(6), 482-491. <https://doi.org/10.1017/s1473550420000269>
- Watson, D., Clark, L. A., & Tellegen, A. (1988). Development and validation of brief measures of positive and negative affect: The PANAS scales. *Journal of Personality and Social Psychology, 54*(6), 1063-1070. <https://doi.org/10.1037/0022-3514.54.6.1063>
- Wiglusz, M., Landowski, J., & Cabała, W. (2019). Psychometric properties and diagnostic utility of the State–Trait Anxiety Inventory in epilepsy with and without comorbid anxiety disorder. *Epilepsy & Behavior, 92*, 221-225. <https://doi.org/10.1016/j.yebeh.2019.01.005>
- Williams, C., Denovan, A., Drinkwater, K., & Dagnall, N. (2021). Thinking style and paranormal belief: the role of cognitive biases. *Imagination, Cognition and Personality, 47*(3), 274-298. <https://doi.org/10.1177/02762366211036435>
- Zimmermann, F., & Kohring, M. (2020). Mistrust, Disinforming News, and Vote Choice: A Panel Survey on the Origins and Consequences of Believing Disinformation in the 2017 German Parliamentary Election. *Political Communication, 37*(2), 215-237. <https://doi.org/10.1080/10584609.2019.1686095>